BAYESIAN PURSUIT ALGORITHM FOR SPARSE REPRESENTATION

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ABSTRACT

In this paper, we propose a Bayesian Pursuit algorithm for sparse representation. It uses both the simplicity of the pursuit algorithms and optimal Bayesian framework to determine active atoms in sparse representation of a signal. We show that using Bayesian Hypothesis testing to determine the active atoms from the correlations leads to an efficient activity measure. Simulation results show that our suggested algorithm has better performance among the algorithms which have been implemented in our simulations in most of the cases.

Index Terms— Sparse representation, Sparse Component Analysis (SCA), Compressed Sensing (CS), Pursuit algorithms.

1. INTRODUCTION

Finding (sufficiently) sparse solutions of underdetermined systems of linear equations (possibly in the noisy case) has been used extensively in signal processing community. This problem has been found applications in a wide range of diverse fields. Some applications are Blind Source Separation (BSS) and Sparse Component Analysis (SCA) [1], decoding [2] and Compressive Sensing (CS) [3], [4].

The problem can be defined in various contexts such as sparse representation, SCA or Compressed Sensing (CS). Here, we use the notation of the sparse representations of signals. Let the model be:

$$\mathbf{x} = \mathbf{\Phi}\mathbf{y} + \mathbf{e} \qquad (1)$$

where $\mathbf{x}$ is an $n \times 1$ signal vector, $\mathbf{y}$ is an $m \times 1$ sparse coefficient vector, $\mathbf{\Phi}$ is an $n \times m$ matrix called dictionary and $\mathbf{e}$ is a $n \times 1$ error vector. It is assumed that $n < m$ which means that the signal length is smaller than the number of columns of the dictionary (which are called atoms). The main assumption is that the signal has a sparse representation in the dictionary. The main goal is to find the sparse coefficient vector $\mathbf{y}$ based on the signal $\mathbf{x}$ and knowing the dictionary $\mathbf{\Phi}$.

In different applications, the interpretations of vectors are different, but in all of them the model is as (1). For example, in the context of CS, $\mathbf{\Phi}$ is the measurement matrix, $\mathbf{x}$ is the very few measurements of the signal and $\mathbf{y}$ is the sparse representation of the true signal in a probable domain. In the context of SCA, $\mathbf{\Phi}$ is the mixing matrix, $\mathbf{x}$ is the mixture vector and $\mathbf{y}$ is the source vector.

Finding the sparsest solution, that is, the solution with the minimum number of nonzero elements, is an NP-hard combinatorial problem. So, different methods have been proposed to solve the problem in a tractable way. Most of them are divided in two main categories: 1) Optimization approaches and 2) Greedy approaches. The first category converts the problem to an optimization problem and then use different methods to solve that. But, the second category tries to find active coefficients (with nonzero elements) directly by an algorithm.

In the first category, the most successful approach which is Basis Pursuit (BP) [5], suggests a convexification of the problem by replacing the ℓ^0-norm with the ℓ^1-norm. It can then be implemented by Linear Programming (LP) methods. Another method is FOCUSS algorithm which uses ℓ^p-norm with $p \leq 1$ as a replacement for the ℓ^0-norm [6]. Recently, a novel Expectation-Maximization (EM) algorithm [7] and a Bayesian Compressive Sensing (BCS) algorithm [8] are proposed to solve the problem in a Bayesian framework. There is also a new method for minimizing a smoothed version of the ℓ^0-norm which is called SL0 method [9].

The methods of the other category choose active coefficients by iterative algorithms. Generally, they use the correlation between the signal (or residual signal) and the atoms of the dictionary as an informative measure for deciding which coefficients are actually active (or nonzero). These algorithms are Matching Pursuit (MP) [10], Orthogonal Matching Pursuit (OMP) [11], Stage-wise OMP (StOMP) [12] and Gradient Pursuit (GP) [13].

Our proposed method, which can be seen as a modification in the Iterative Detection-Estimation (IDE) algorithm [14], uses the simplicity of the greedy pursuit algorithms while simultaneously uses Bayesian tools for optimal selection of active atoms.

2. SYSTEM MODEL

The noise vector in the model (1) is assumed to be zero-mean Gaussian with covariance matrix $\sigma_e^2 \mathbf{I}$. We model the sparse coefficients as follows. In our model the coefficients are inactive with probability p, and are active with probability $1 - p$ (sparsity of $\mathbf{s}$ implies that p should be near 1). In the inactive case, the values of the coefficients are zero and in the active case the values are obtained from a Gaussian distribution. We call this model the 'spiky model' which is a special case of the Bernoulli-Gaussian model with the variance of the inactive samples being zero. This model has been also used in [7]. It is suitable for sparse representation of a signal where we want to decompose a signal as a combination of only a few atoms of the dictionary and the coefficients of the other atoms are zero. So, the probability density of the coefficients in our problem is:

$$p(y_i) = p\delta(y_i) + (1 - p)N(0, \sigma_r^2) \qquad (2)$$

In this model, each coefficient can be written as $y_i = q_i r_i$ where q_i is a binary variable (with binomial distribution) and r_i is the amplitude of the i'th coefficient with Gaussian distribution. Each element q_i shows the activity of the corresponding coefficient (or corresponding atom). That is:

$$q_i = \begin{cases} 1 & \text{if } y_i \text{ is active} \quad (\text{with probability } 1 - p) \\ 0 & \text{if } y_i \text{ is inactive} \quad (\text{with probability } p) \end{cases} \qquad (3)$$

Consequently, the probability $p(\mathbf{q})$ of activity vector $\mathbf{q} \triangleq (q_1, q_2, ..., q_m)^T$ is equal to:

$$p(\mathbf{q}) = (1 - p)^{n_a}(p)^{m - n_a} \qquad (4)$$

where n_a is the number of active coefficients, i.e., the number of 1's in $\mathbf{q}$. So, the coefficient vector can be written as:

$$\mathbf{y} = \mathbf{Qr}, \qquad \mathbf{Q} = \text{diag}(\mathbf{q}) \qquad (5)$$

where $\mathbf{q}$ and $\mathbf{r} \triangleq (r_1, r_2, ..., r_m)^T$ are the 'activity vector' and 'amplitude vector', respectively.

3. BAYESIAN PURSUIT ALGORITHM

The main task in sparse recovery algorithms is to determine which atoms are active in the sparse representation of the signal. In some pursuit algorithms (e.g., MP), it is determined by correlation maximization. In some other pursuit algorithms (e.g., StOMP), it is done by comparing the correlations with a threshold. But, here we want to determine it by a Bayesian hypothesis testing from the correlations. We will see that the same activity measure as the IDE algorithm [14] is obtained with the difference that the threshold is obtained mathematically. To develop the hypothesis testing in our Bayesian Pursuit Algorithm (BPA), we write (1) as:

$$\mathbf{x} = \sum_{i=1}^{m} \boldsymbol{\varphi}_i y_i + \mathbf{e} \qquad (6)$$

where $\boldsymbol{\varphi}_i$ is the columns of the dictionary or atoms. So, the correlations between the original signal and the atoms are:

$$z_j \triangleq <\mathbf{x}, \boldsymbol{\varphi}_j> = y_j + \sum_{i \neq j} y_i b_{ij} + v_j \qquad (7)$$

where $b_{ij} \triangleq <\boldsymbol{\varphi}_i, \boldsymbol{\varphi}_j>$ and $v_j \triangleq <\mathbf{e}, \boldsymbol{\varphi}_j>$ and the atoms are assumed to have unit norm.

To do a Bayesian hypothesis testing based on correlations for determining the activity of the j'th atom, we should compute the posteriors $p(H_1|\mathbf{z})$ and $p(H_2|\mathbf{z})$, where the hypothesis H_1 is the hypothesis that the j'th atom is active and H_2 is the hypothesis that the j'th atom is inactive. To obtain a simple algorithm like pursuit algorithms, we assume that we know the previous estimations of other coefficients (except the j'th coefficient). And now, we want to know the activity of only the j'th atom and then update just only the j'th coefficient.

Since we assume that we know the previous estimations of other coefficients, (7) can be written as:

$$z_j - \sum_{i \neq j} \hat{y}_i b_{ij} = y_j + \sum_{i \neq j} (y_i - \hat{y}_i) b_{ij} + v_j \qquad (8)$$

where $\hat{y}_i$ is the estimation of the i'th coefficient up to the current iteration. With the following definitions:

$$m_j \triangleq \sum_{i \neq j} \hat{y}_i b_{ij}$$

$$\gamma_j \triangleq \sum_{i \neq j} (y_i - \hat{y}_i) b_{ij} + v_j \qquad (9)$$

The two hypothesis H_1 and H_2 are:

$$\text{Hypotheses}: \begin{cases} H_1 : z_j - m_j = r_j + \gamma_j \\ H_2 : z_j - m_j = \gamma_j \end{cases} \qquad (10)$$

where m_j is known and γ_j has a flavour of noise and error. (10) resembles a classical detection problem. The optimal hypothesis testing involves the computation of the overall posteriors $p(H_1|\mathbf{z})$ and $p(H_2|\mathbf{z})$. But, with the previous assumptions and formulations, we reach a relatively simple detection problem as in (10). So, for the simplicity of the algorithm like the pursuit algorithms, we rely only on the simpler posteriors as $p(H_1|z_j)$ and $p(H_2|z_j)$. So, the hypothesis H_1 is assumed to be true when $p(H_1|z_j) > p(H_2|z_j)$. Based on the Bayes rule, the above posteriors are proportional to $p(H_1)p(z_j|H_1)$ and $p(H_2)p(z_j|H_2)$ respectively. The prior probabilities for the hypotheses are $p(H_1) = 1 - p$ and $p(H_1) = p$ where p is defined in Section 2. We assume that the coefficient errors $(y_i - \hat{y}_i)$ have a Gaussian distribution with variance σ_{i,e_y}^2. Hence, by assuming that the error $(y_i - \hat{y}_i)$ is Gaussian, the term γ_j is Gaussian and we assume its variance is $\sigma_{\gamma_j}^2$. Therefore, we have:

$$\frac{(1 - p)}{\sqrt{2\pi(\sigma_{\gamma_j}^2 + \sigma_r^2)}} \exp\left(\frac{-(z_j - m_j)^2}{2(\sigma_{\gamma_j}^2 + \sigma_r^2)}\right) >$$

$$\frac{p}{\sqrt{2\pi\sigma_{\gamma_j}^2}}\exp(\frac{-(z_j-m_j)^2}{2\sigma_{\gamma_j}^2}) \qquad (11)$$

Simplifying (11) with the assumption that the unknown parameters (p, σ_r and σ_{γ_j}) are known, leads to the following decision rule for the hypothesis testing:

$$\text{Activity}(y_j) \triangleq |z_j-m_j| > \text{Th}_j \qquad (12)$$

where Th_j is defined as:

$$\text{Th}_j \triangleq \frac{\sigma_{\gamma_j}}{\sigma_r}\sqrt{2(\sigma_r^2+\sigma_{\gamma_j}^2)\ln(\frac{p}{1-p}\frac{\sqrt{\sigma_r^2+\sigma_{\gamma_j}^2}}{\sigma_{\gamma_j}})} \qquad (13)$$

Although (13) determines the optimum threshold, it depends on unknown parameters (p, σ_r and σ_γ) which should be estimated from the original signal ($\mathbf{x}$). To estimate the parameters p, σ_r and σ_e, we can use similar formulas as in [7] which are:

$$\hat{p}=\frac{||\mathbf{q}||_0}{m}, \hat{\sigma_e}=\frac{||\mathbf{x}-\mathbf{\Phi}\hat{\mathbf{y}}||_2}{\sqrt{n}}, \hat{\sigma_r}=\frac{||\mathbf{r}||_2}{\sqrt{m}} \qquad (14)$$

The problem here is to estimate the parameter σ_{γ_j} which is the standard deviation of γ_j in (9). We assume the independence between v_j and coefficient error ($y_i-\hat{y}_i$). Another assumption is the independence between distinct coefficient errors $y_i-\hat{y}_i$ and $y_j-\hat{y}_j$ for $i\neq j$. It is also known that v_j is a Gaussian random variable with the similar variance as e_j which is uniformly σ_e^2. So, we have the following formula for the parameter estimation:

$$\sigma_{\gamma_j}^2 = \sigma_e^2 + \sum_{i\neq j} b_{ij}^2\sigma_{i,e_y}^2 \qquad (15)$$

where σ_{i,e_y}^2 is the variance of the coefficient error $y_i-\hat{y}_i$. If our algorithm converges, we expect that σ_{i,e_y} decreases. So, we force that this error variance decreases linearly with a coefficient α which is less but near one. So, we select this decreasing sequence as:

$$\sigma_{i,e_y}^{(n+1)} = \alpha\sigma_{i,e_y}^{(n)} \qquad (16)$$

where parameter α determines the rate of convergence.

As we can see from (13), the optimum threshold is changed from an initial large value to a small final value. The initial and final values of the threshold are:

$$\text{Th}_j^{(0)} = \text{Th}|_{\sigma_{\gamma_j}^{(0)}}$$

$$\text{Th}_j^{(\infty)} = \text{Th}|_{\sigma_\gamma=\sigma_e} \approx K\sigma_e \qquad (17)$$

where $K=\sqrt{2\ln(\frac{p}{1-p}\frac{\sigma_r}{\sigma_e})}$. As we can see from (17), the initial thresholds are different from one coefficient to another. But, all the thresholds have converged to the same threshold which does not depend on the coefficient.

As the value of threshold changes from a large value to a small value, the algorithm can detect more and more atoms. At first iterations, the optimal thresholding strategy in (13) changes the thresholds very fast and then after some iterations, the thresholds converge to the final small value.

After updating the activity vector based on BPA decision rule in (12), the estimation of amplitude vector $\mathbf{r}$ which was defined in Section 2, based on this estimated activity vector can be done with a Linear Least Square (LLS) estimation [7] as:

$$\hat{\mathbf{r}} = \sigma_r^2\hat{\mathbf{Q}}\mathbf{\Phi}^T(\sigma_r^2\mathbf{\Phi}\hat{\mathbf{Q}}\mathbf{\Phi}^T+\sigma_e^2\mathbf{I})^{-1}\mathbf{x} \qquad (18)$$

where $\hat{\mathbf{Q}}=\text{diag}(\hat{\mathbf{q}})$ and $\hat{\mathbf{q}}$ is the updated activity vector.

4. EXPERIMENTS

The performance of the proposed BPA algorithm is investigated in this section. The comparison of our BPA algorithm is done with some other algorithms in the literature in both estimation accuracy and complexity viewpoints. The estimation accuracy of the algorithms are compared with the Signal to Noise Ratio between the true coefficients and the recovered coefficients, which is defined as:

$$\text{SNR}_o \triangleq 10\log(\frac{||\mathbf{y}||_2^2}{||\mathbf{y}-\hat{\mathbf{y}}||_2^2}) \qquad (19)$$

where the index indicates that it is an output SNR. We define another measure which determines the noise level. We refer to it as input SNR and is defined as:

$$\text{SNR}_i \triangleq 20\log(\frac{\sigma_r}{\sigma_e}) \qquad (20)$$

We use the CPU time as a measure of complexity. Our simulations were performed in MATLAB7.0 environment using an AMD Athlon Dual core 4600 with 896 MB of RAM and under Windows Xp operating system.

In our experiment, we used a random dictionary matrix with uniform distributed elements from $[-1,1]$, and then normalized its columns. The dimension of our problem was selected as $m=512$ for the number of atoms and $n=256$ for the signal length. For generating the sparse coefficients, we used the model in (2) with the probability $p=0.9$ and unit variance for the active coefficients ($\sigma_r=1$). So, approximately 51 atoms are active in the sparse representation of the signal. The noise level or error is considered to have a Gaussian distribution with different variances. The measure of performance which is the output SNR in (19) is averaged over 100 different random realizations of the dictionary, sparse coefficients and noise vector.

For initializing the unknown statistical parameters (p, σ_r and σ_e), we use $\hat{p}^{(0)}=0.8$, $\hat{\sigma_r}^{(0)}=\frac{||\mathbf{x}||_2}{\sqrt{m(1-\hat{p}^{(0)})}}$ and $\hat{\sigma_e}^{(0)}=\frac{\hat{\sigma_r}^{(0)}}{5}$ which are similar to those used in [7]. The important note is to use an overestimate of noise variance (σ_e) in the initial
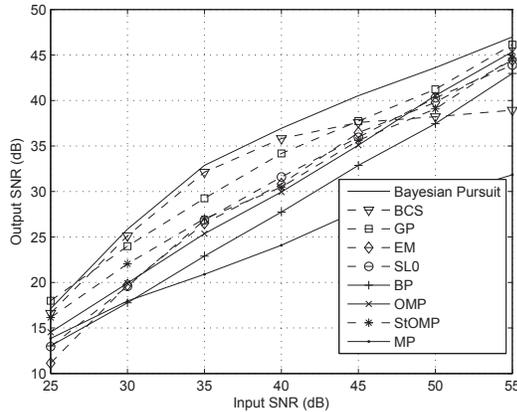
Fig. 1. The output SNR versus input SNR for various algorithms. The parameters are $m = 512$, $n = 256$, $p = 0.9$, $\sigma_r = 1$, $\alpha = 0.9$ and 20 iterations are used for BPA.

iteration. We also used 20 iterations for BPA algorithm and the simulation parameter α was selected as 0.9.

In this experiment, we compared the suggested BPA algorithm with BP, MP, OMP, StOMP, SL0, EM and GP. For MP and OMP, we used 100 and 50 iterations, respectively (We used codes from Sparse Lab toolbox http://sparselab.stanford .edu/). For StOMP, we used 20 iterations and the sensitivity parameter for threshold selection was selected 0.5 (refer to Sparse Lab toolbox http://sparselab.stanford.edu/). For SL0, we used the minimum sigma equal to 0.04 and the decreasing factor equal to 0.9 (refer to http://ee.sharif.edu/ SLzero/sl0.m). For the EM algorithm, we used both 5 iterations for the overall EM algorithm and 4 iterations for the M-step [7]. Figure 1 shows the estimation accuracy of various algorithms versus the noise level. It shows that our BPA algorithm outperforms the others in most of the cases.

We computed the average simulation time for various algorithms. These are 0.0083, 0.0922, 0.0049, .0800, 0.4737, 0.0172, 0.6412, 0.6313 and 0.0207 seconds for MP, SL0, GP, BCS, BP, StOMP, BPA, EM and OMP respectively. So, the BPA algorithm is the most complex algorithm.

5. CONCLUSION

In this paper, we suggested the BPA algorithm which determines the active atoms based on a Bayesian hypothesis testing from the correlations of the signal with the atoms of the dictionary. Simulations show the advantage of the proposed method over some of the state-of-the-art algorithms in terms of estimation accuracy in most of the cases.

6. REFERENCES

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